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### ► To cite this version:

Beth A Davison, Stefanie Senger, Iziah Sama, Gary G Koch, Alexandre Mebazaa, et al.. Is acute heart failure a distinctive disorder? An analysis from BIOSTAT-CHF. European Journal of Heart Failure, 2021, 23 (1), pp.43-57. 10.1002/ejhf.2077 . hal-03759089

## HAL Id: hal-03759089 https://hal.univ-lorraine.fr/hal-03759089v1

Submitted on 23 Aug 2022

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### Is Acute heart failure a distinctive disorder? An analysis from BIOSTAT-CHF

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#### Total word count: 3895

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#### Abstract

*Aims:* This retrospective analysis sought to identify markers that might distinguish between acute heart failure (HF) and worsening HF in chronic outpatients.

*Methods and Results:* The BIOSTAT-CHF index cohort included 2516 patients with new or worsening HF symptoms: 1694 enrolled as inpatients (acute HF) and 822 as outpatients (worsening HF in chronic outpatients). A validation cohort included 935 inpatients and 803 outpatients. Multivariable models were developed in the index cohort using clinical characteristics, routine laboratory values, and proteomics data to examine which factors predict adverse outcomes in both conditions and to determine which factors differ between acute HF and worsening HF in chronic outpatients, validated in the validation cohort.

Patients with acute HF had substantially higher morbidity and mortality (6 months mortality was 12..3% for acute HF and 4..7% for worsening HF in chronic outpatients). Multivariable models predicting 180-day mortality and 180-day HF re-admission differed substantially between acute HF and worsening HF in chronic outpatients. CA-125 was the strongest single biomarker to distinguish acute HF from worsening HF in chronic outpatients, but only yielded a C-index of 0..71. A model including multiple biomarkers and clinical variables achieved a high degree of discrimination with a C-index of 0..913 in the index cohort and 0..901 in the validation cohort. *Conclusion:* The study identifies different characteristics and predictors of outcome in acute HF patients as compared to outpatients with chronic HF developing worsening HF. The markers identified may be useful in better diagnosing acute HF and may become targets for treatment development.

#### **Keywords**

Acute heart failure; acute heart failure diagnosis; acute heart failure treatment.

#### Introduction

Acute heart failure (AHF) is defined as a worsening of symptoms and signs of heart failure (HF) requiring urgent care inclusive (but not limited to) intravenous therapy and hospital admission.(1) Following a series of neutral studies in which new interventions for AHF were not shown in large studies to be associated with improvements in either patients' symptoms or shortand long-term outcomes,(2) some opinion leaders have raised doubt whether AHF is a separate condition or just a part of the natural course of chronic heart failure.(3) At the same time doubts have been raised that some of the failure in demonstrating positive effects in large AHF studies relates to dilution of the patient population;(4) i.e., that patients enrolled in larger confirmatory studies may not have "true" AHF but have other disorders, chiefly chronic HF that has slightly deteriorated.(4) On the one hand some argue that the decision to admit a patient with AHF is subjective and variable and hence HF deterioration managed in the outpatient setting is not a different entity than AHF leading to hospital admission, while others argue that those are different conditions, the latter having a distinct pathophysiological mechanism – associated with inflammatory activation, more congestion, and end organ damage.(5) One of the obstacles in resolving these differences, and possibly developing effective therapies for AHF, is that there are no objective measures that help determine whether a patient is truly "acute", i.e., has AHF. All objective measures utilized in the diagnosis of AHF to date (natriuretic peptides levels, chest Xray or lung ultrasound) can also be found to be affected in patients with stable chronic HF who have slight outpatient deterioration.

In the current analysis we have examined the characteristics of patients enrolled in the BIOSTAT-CHF study (A systems BIOlogy Study to Tailored Treatment in Chronic Heart

Failure) study where patients with both AHF requiring hospital admission and worsening of HF managed in an outpatient clinic were enrolled and followed.(6) While the absolute risk of adverse clinical outcomes overlapped in these two patient groups, the characteristics and prognosis of patients enrolled in the inpatient versus the outpatient setting in this study were found to differ.(7) The objective of this retrospective analysis was to identify markers that may distinguish between AHF and chronic HF patients with outpatient exacerbations and determine whether predictors of adverse outcomes in the two groups differed in the BIOSTAT-CHF database.

#### Methods

#### Index and validation cohorts

The BIOSTAT-CHF project provided access to data from an index cohort of 2516 heart failure patients enrolled between December 2010 and December 2012 in 69 centers in 11 European countries. In the Index Cohort adult patients with new or worsening heart failure symptoms, objective evidence of cardiac dysfunction, treated with at least 40 mg/day furosemide or equivalent, and receiving 50% or less the target doses of evidence-based therapies were enrolled in either the inpatient or outpatient setting and patients were followed for a median of 21 months.(6). For the Validation cohort inclusion criteria were similar, although outpatients could have been recruited without worsening of heart failure. Data were available from a validation cohort of 1738 patients enrolled in either the inpatient or outpatient setting and patients in the validation cohort had a heart failure diagnosis and a previous admission with heart failure requiring diuretic therapy, and were treated with at least 20 mg/day furosemide or equivalent and 50% or less the target doses of

evidence-based therapies at entry. The study design of BIOSTAT-HF, baseline characteristics of the two cohorts and various modeling results have been described previously (6-8).

#### Patient characteristics and clinical outcomes

In order to develop a discriminatory model for acute HF and outpatient worsening HF (OP-WHF) we used reported hospitalization status (inpatient or outpatient) as surrogate. Model development was based on basic patient characteristics collected at study entry and baseline (age, sex, LVEF, BMI, vital signs, medical history), as well as baseline laboratory test values from blood samples analyzed locally or from blood samples that were frozen for shipment to central laboratories for analysis:

- Local and routine central laboratory: Local laboratory results considered in the models included white blood cell count, red blood cell count, platelet count, hemoglobin, urea, glomerular filtration rate estimated from local creatinine using the CKD-EPI equation,(9) sodium, potassium, total bilirubin, glucose, AST, ALT, GGT, ALP, HDL, LDL, and triglycerides. NT-proBNP, high-sensitivity troponin T (hsTnT), and GDF-15 were centrally measured using a Roche Elecsys® cobas analyzer (Roche Diagnostics, Mannheim, Germany); aldosterone, renin, FGF23, urea, creatinine, calcium, phosphate, albumin, iron, ferritin, transferrin, hepcidin, and sTFR were also measured centrally.
- Specialty biomarkers: Three biomarkers (TnI, ET-1, and IL-6) were measured using enzyme-linked immunosorbent assays (Singulex Inc.) on a Luminex platform. pro-ENK and bio-ADM were measured on a Sphingotec platform (Spingotec GmbH). CA125 was measured using a chemiluminescent microparticle immunoassay on an ARCHITECTi system (Abbott Laboratories).

3. Proteinomics/Olink panels: The BIOSTAT-CHF project included a comprehensive proteomic database measured by the Olink Proseek analysis service (Olink Proteomics, Uppsala, Sweden). The Olink platform utilizes a high-throughput multiplex immunoassay based on a proprietary Proximity Extension Assay (PEA) technology, where each biomarker is addressed by a matched pair of antibodies, coupled to unique, partially complementary oligonucleotides, and measured by quantitative real-time polymerase chain reaction (PCR). Results are expressed in the form of relative quantification (Normalized Protein eXpression or NPX) which are logarithmically related to protein concentration but cannot be converted to absolute protein concentrations. The BIOSTAT-CHF database comprised data from four Olink panels: Cardiovascular II, Cardiovascular III, Immune Response, and Oncology II, providing baseline measurements for a total of 368 proteins for most patients of both cohorts. Proteins can be recognized by a UniProt identifier.(10) For model development, we excluded 4 biomarkers from analysis as the same proteins were measured on more than 1 panel (1 copy each of AREG and SCF, and two copies of IL-6), leaving a total of 364 unique proteins for analysis.

In a preliminary step, we excluded baseline parameters with mostly missing observations in the index cohort and selected representative parameters in case of highly correlated or collinear variables. Olink measures were nearly perfectly correlated with other central measures of the same parameter. When both were available models were developed considering only the Olink parameter; for example, only MUC16 and not CA125 was considered in multivariable Model 3 for distinguishing AHF from OP-WHF. A differential expression analyses of the Olink proteins was performed using the Linear Models for Microarray data analysis (Limma) software (version 3.34.9).(11) Proteins were considered differentially expressed in inpatients relative to outpatients

if the absolute value of the fold change exceeded 1.24 ( $|\log_2 FC|>0.31$ ), the p-value for the t-test that the  $\log_2FC$  differs from zero and the false discovery rate (FDR) <0.05. A volcano plot [a plot of  $-\log_{10}(p-value)$  versus  $\log_2$  fold change] was used to visualize the differential expression of these markers in patients enrolled in the inpatient versus the outpatient setting. Clinical endpoints considered for analysis were the two components of the primary endpoint of BIOSTAT-CHF: all-cause death and first re-admission for heart failure. For the development of prognostic models, the time to the first occurrence of each clinical endpoint was evaluated through 180 days after baseline. Partially missing re-admission or death dates were imputed with the 15<sup>th</sup> of the month if only day was missing, or July 1<sup>st</sup> if both day and month were missing. Re-admission dates before and up to ( $\leq$ ) the respective baseline visit dates were excluded from analysis. Time to event was computed for all patients with a recorded event. Time to event for patients without a recorded event was censored at the earlier of 180 days after baseline or the individual end of study date. For the analysis of time to first occurrence of HF re-admission, time was censored at date of death for subjects who were not re-admitted for HF as the primary cause.

#### Statistical model development

Hospitalization status (inpatient versus outpatient) was analyzed using logistic regression. Time to all-cause death through 180 days and time to first HF re-admission through 180 days were analyzed using Cox proportional hazards models. We examined a number of classification methods in addition to logistic and Cox regression including Boosted Logistic Regression (LogitBoost), linear discriminant analysis with stepwise feature selection (stepLDA), neural networks (nnet), k-nearest neighbors (knn), CART (rpart), C5.0, and random forest (rf) (Supplementary Tables 1-5). We noted that performance characteristics of C5.0 and random

forest appeared best, with generally the highest area under the receiver operator characteristic curves (AUCs) and highest scaled Brier scores (which can be interpreted as correlation coefficients). Thus, pre-selection of Olink candidate predictors for the logistic regression models was based on the average variable importance rank for these two methods. We used the set of index cohort patients with complete Olink data as a training set for building classifiers that would predict inpatient status. Variable importance for each Olink marker was calculated by independently using C5.0(12) as well as Random Forests(13) as classification methods in a repeated 5-fold cross-validation approach using R package "caret"(14). Variable importance measures within each method were ranked using sports ranking; markers were then sorted by their average rank across the two methods. The top 50 Olink markers were selected as candidates for the inpatient status model (Supplementary Table 6), and the top 20 markers as candidates for the outcome models except for the top 10 in the case of outpatient mortality.

The full set of candidate predictors for each of the five models is given in Supplementary Table 11. For discriminating between in- and outpatients, in Model 1 we considered only patient characteristics, vital signs and locally- and centrally-measured laboratory values; we additionally considered medical history for prognostic models. For model 2, we further considered specialty biomarker measures; for model 3, we added the pre-selected Olink proteins. Missing values in the final analysis data set were imputed using a multi-chain Monte Carlo approach (R package "mice"(15)) and 10 imputed data sets were generated for the index cohort.

For each model, each continuous candidate baseline variable was first tested for non-linearity of its association with the model outcome by assessing the significance of the non-linear components of a restricted cubic spline transformation applied to the baseline variable in the index cohort while adjusting for the remaining candidate variables. In cases where the

association was deemed significantly non-linear and this behavior was observed consistently across the 10 imputed data sets of the index cohort, appropriate non-linear transformations were selected from a set of pre-specified transformations (such as quadratic, cubic, or linear spline transformations). Selection was based on values of Akaike's Information Criterion and visual inspection of plotting the predicted outcome against the baseline values. Baseline variables with highly skewed distributions were log2 transformed for analysis. Logistic or Cox regression with backwards selection was run on the 10 imputed data sets computing pooled p-values according to Rubin's algorithm(16) in each selection step and with the p-value criterion of 0.01 for staying in the inpatient models, and 0.05 for staying in the prognostic models. Estimated effect sizes, their 95% confidence intervals and p-values were pooled using Rubin's algorithm.

As a measure of discriminatory ability, the C-index pooled across the 10 imputed data sets was computed for each model. We further derived "final" models by combining the regression coefficients using Rubin's algorithm, applied each final model on the imputed data sets and derived the C-index for the final models as the average C-index across the imputed data sets. For internal validation, bootstrap samples of the imputed data sets of the index cohort were drawn, each time using the same random sample of patients for all imputations. Backwards selection and model fitting were repeated for each bootstrap sample in order to estimate bias-corrected C-indices and confidence intervals for each multivariable model. For external validation, the final models were applied to the imputed data sets of the validation cohort and pooled C-indices estimated. We further examined discrimination and calibration of our models through receiver-operator characteristics (ROC) curves and calibration plots.

The model for inpatient versus outpatient status was externally validated in the validation cohort. A few of the baseline characteristics included in the final models were not accessible or not

reported in the validation cohort and were thus multiply imputed. We successively added a small random sample of validation patients (n=11) to each of the 10 imputed data sets of the index cohort. We then imputed missing data for each subset of 11 validation patients in each imputed data set, repeated this step 158 times, thus generating 10 imputed data sets for the 1738 subjects in the validation cohort as well.

To further explore differences in prognostic factors between inpatients and outpatients, we applied the final multivariable models in inpatients to the outpatients. For HF readmission, all variables in the final model were used, while for death due to the limited number of events in outpatients the top 8 predictors were chosen. We further compared the fit of the final inpatient and outpatient models in the inpatients using partial likelihood ratio tests(17) as implemented in R package nonnestcox; for each clinical endpoint, the two final models were fitted within the inpatients and separately for 10 imputed data sets.

SAS® version 9.4 (SAS Institute, Cary, NC, USA) and R version 3.5.1(18) software was used for all analyses.

#### Data availability

The data that support the findings of this study are available from University Medical Center Groningen but restrictions apply to the availability of these data, which were used under license for the current study, and so are not publicly available. Data are however available from the authors upon reasonable request and with permission of University Medical Center Groningen.

#### Results

Inpatients versus outpatients

Summary statistics for selected baseline characteristics are presented in Table 1 by cohort and patient status. A volcano plot showing the differential expression of the Olink panel proteins in inpatients versus outpatients is presented in Figure 1. Mucin-16 (MUC16), also known as CA125, appears to be the most differentially expressed, when considered without simultaneous adjustment for other proteins (i.e., univariably). In BIOSTAT-CHF, a doubling of CA125 (ARCHITECTi) as a continuous measure increased the odds of being an inpatient by 50% (OR 1.50, 95% CI 1.42-1.58, p<0.0001) with an AUC of 0.6983 (Table 2); results for MUC-16 measured on the Olink platform were similar (OR 1.58, 95% CI 1.48-1.69, p<0.0001). MUC-16 was selected for inclusion in the multivariable model (Table 3) with a somewhat smaller association with inpatient status (OR 1.30, 95% CI 1.17-1.46). At a threshold of 100 U/mL,(19) CA125 (ARCHITECTi) greater than the threshold had an AUC of 0.6278, sensitivity of 0.3832, specificity 0.8725, positive predictive value (PPV) 0.8610 and negative PV (NPV) 0.4070. ..... We also examined the ability of the traditional marker of HF severity - NT-proBNP - to discriminate between inpatients and outpatients. A doubling of NT-proBNP (cobas) was associated with an OR of 1.32 (95% CI 1.25-1.39, p<0.0001) for inpatient status with an AUC of 0.6395. Using a cut-point of 400 pg/mL(20) resulted in an AUC of 0.5421, sensitivity 0.9400, specificity 0.1442, PPV 0.6936, and NPV 0.5384. Thus, the traditional cut-point displayed high sensitivity and very low specificity. ......Note that when adjusted simultaneously for other proteins (Table 3), patients with a higher NT-proBNP were less likely to be an in-patient (OR 0.87 per 1-log increase, 95% CI 0.78-0.96).

We examined the ability of other biomarkers identified as potential markers of acutely ill heart failure patients including ST2, troponin T, troponin I, GDF-15, and ADM. Each of these markers

individually was unable to discriminate between inpatients and outpatients, with AUCs of about 0.500 signifying an ability no better than chance to predict the status.

Table 3 presents the selected multivariable logistic regression models for inpatient versus outpatient; Supplementary Figures 1 and 2 show in the index and validation cohorts the receiver-operator-characteristic (ROC) and calibration curves, respectively. The discrimination of the final multivariable model 3, which included 21 of the 50 Olink proteins considered, was excellent, with a c-index (AUC) of 0.9133 in the index cohort and of 0.9011 in the validation cohort. The performance characteristics of this model – including AUC and scaled Brier score – were better than those using classification methods (Supplementary Table 1).

#### Death through Day 180

In the Index cohort a total of 208 (12.3%) patients enrolled in the inpatient setting, and 39 (4.7%) patients enrolled in the outpatient setting, died by day 180. In the validation cohort 162 (17.3%) of the patients enrolled in the inpatient setting died at 6 months versus 25 (3.1%) in the outpatient setting. Final multivariable prognostic models are presented in Tables 4 and 5 for inpatients and outpatients, respectively. The full models (Model 3) had good discrimination with c-indexes of 0.8281 and 0.8460, respectively. Although difficult to compare directly because the paucity of events among outpatients restricted the number of candidate predictors that could be considered for that outcome, the prognostic factors differed for the two patient groups. When only considering patient characteristics and local and central laboratory data (Model 1), for example, FGF23, NT-proBNP, renin, and troponin T were all highly prognostic in inpatients, while platelet count, peripheral arterial disease, and age were most prognostic in outpatients.

The 8 most prognostic factors from the inpatient multivariable Model 3 for death provided much less discrimination in outpatients with a c-index 0.7464 (Supplementary table 12). And the inpatient and outpatient models were found to be distinguishable, and the fit of the inpatient model significantly better than the outpatient model, in inpatients.

#### Heart failure hospitalization through Day 180

In the Index Cohort, a total of 254 (15.0%) patients enrolled as inpatients, and 73 (8.9%) patients enrolled as outpatients, were hospitalized for heart failure by day 180. In the Validation Cohort, 6-month HF admission was observed in 166 (17.8%) inpatients and 57 (7.1%) outpatients. Final multivariable prognostic models are presented in Tables 6 and 7 for inpatients and outpatients, respectively. Model discrimination was modest in inpatients, with a c-index of 0.7322 for the full model (Model 3); the model including only patient characteristics and local or central laboratory data (Model 1) and the model additionally considering specialty laboratory parameters (Model 2) had similar discrimination with a c-index of 0.7395 for both. Discrimination for the models in outpatients was better, with c-indexes of 0.7966, 0.7984, and 0.8234 for Models 1, 2, and 3, respectively. NT-proBNP was a strong prognostic factor in both inpatients and outpatients. The inpatient Model 3 applied to outpatients provided less discrimination than the model developed in the outpatients, with a c-index of 0.8055 (Supplementary table 13). And the inpatient and outpatient models were found to be distinguishable, and the fit of the inpatient model significantly better than the outpatient model, in inpatients.

#### Discussion

AHF research has been limited in the last decades by three major issues. The first is the lack of an objective definition of AHF. The currently used definition is a subjective one (worsening

symptoms and signs requiring urgent care) which has been an impediment to distinguishing between patients with "true" AHF versus those with out-patient HF exacerbations not requiring urgent care. This had led to significant problems in enrolling AHF patients in large studies.(4) Second, related to the lack of ability to define AHF we also know little of its pathophysiology. And third, as a consequence of our lack of ability to define AHF and our lack of knowledge of its pathophysiology the treatment targets for AHF are also not defined. In the last two decades, most therapies developed for AHF were either vasodilators or diuretics – which have both failed to show substantial benefit beyond some improvement in very short-term symptoms. In the current analysis, AHF and OP-WHF were found to differ in three major domains. First, patients with AHF had much higher morbidity and mortality rates. The 180-day mortality and HF readmission rates were 12.3% and 15% for the AHF cohort and 4.7% and 8.9% for the OP-WHF cohort, respectively. Second, prognostic models for adverse outcomes differ for patients with AHF versus outpatient exacerbation of HF and models that predict adverse outcomes in AHF do not predict well adverse outcomes in outpatient exacerbations of HF. Third and lastly, the characteristics and prognosis of patients enrolled in the inpatient versus the outpatient setting in this study were found to differ.(7) The current analysis suggests that patients admitted for AHF have a different biomarker profile from patients with HF exacerbation not requiring admission.

These findings suggest that different pathophysiological mechanisms leading to different patterns of activation of neurohormonal and inflammatory protein markers may be involved in AHF and differ from those in out-patient exacerbations of heart failure. If confirmed, these may enable development of new diagnostic platforms that would lead to better ability to diagnose patients with true AHF. The exact components and details of such diagnostic platforms should be

elucidated in further prospective studies. In line with this limitation, as seen above, none of the currently proposed biomarkers (natriuretic peptides, CA125, ST2 or troponin) by itself can differentiate between AHF and outpatient exacerbation to the degree that the full model can. However, more data are required to better elucidate which variables and biomarkers would best discriminate the different disorders. As described in the current manuscript the models provided may help illuminate new paths in developing better models, but are not in themselves ready to be applied immediately in clinical studies.

In addition, some biomarkers that are differentially activated in AHF or those that seem to be associated with adverse outcomes in AHF patients specifically may become targets for therapeutic interventions or proxies to therapeutic intervention success enabling more targeted therapy for AHF to be developed. Oncological development plans have for the last 20 years been successful in targeting specific phenotypes with specific tailored therapies targeting the pathways most activated in those phenotypes. Novel approaches assessing in parallel multiple targeted interventions in specific phenotypes should be adopted in AHF research. It is possible that the lack of success we have encountered in developing new therapies for AHF is not related to the proposal that AHF does not exist, but rather to our limited attempts to develop tailored phenotype specific treatments.

#### Study Limitations

The current analysis is limited by the moderate size of the BIOSTAT study which was designed mainly to examine the importance of treatment optimization in patients with worsening HF. Some of the outcomes were sparse – especially in the group of patients with exacerbation of HF not requiring admission –. This was especially true in the Validation Cohort where some patients

could have been enrolled in the outpatient setting without worsening of HF. Moreover, the BIOSTAT study included a cohort of European mostly Caucasian patients mostly with left ventricular systolic dysfunction. Therefore, our models need validation in other cohorts of more diverse patient populations. The prognostic models developed have relatively low predictive value especially when it comes to HF readmissions. Therefore, they cannot be suggested to replace currently validated models in acute and chronic HF.

#### Conclusions

Our analysis suggests that patients who present with AHF differ from patients who develop HF exacerbation not requiring hospital admission. Patients with AHF are characterized by different clinical and biomarker profiles, have substantially worse outcomes and different predictors of adverse outcomes. The biomarkers that differ between patients with AHF and outpatients with HF exacerbation as well as the predictors of adverse outcome in AHF patients can serve to improve AHF diagnosis and potentially become therapeutic targets for AHF.

Funding

BIOSTAT-CHF was funded by the European Commission [FP7-242209-BIOSTAT-CHF; EudraCT 2010-020808-29].

#### **Conflicts of Interest**

B.A.D. is an employee of Momentum Research who received research grants from Abbott Laboratories, Amgen, Celyad, Cirius Therapeutics, Roche Diagnostics, Sanofi, and Windtree Therapeutics Inc.

S.S. is a former employee of Momentum Research who received research grants from Abbott Laboratories, Amgen, Celyad, Cirius Therapeutics, Roche Diagnostics, Sanofi, and Windtree Therapeutics Inc.

I.S. reports no conflicts.

G.G.K. is the principal investigator of a biostatistical agreement between Momentum Research, Inc. and the University of North Carolina at Chapel Hill, and that agreement provided the structure for his activity for this article. He is also the principal investigator of many such biostatistical agreements with other biopharmaceutical sponsors, including AbbVie, Amgen, Arena, AstraZeneca, Eli Lilly & Co., Forest Research Institute (Allergan), GlaxoSmithKline, Merck, Novartis, Otsuka, Pfizer, and Sanofi, although his activities for those sponsors are not related to the content of this article. Information concerning all biostatistical agreements for which Gary Koch is the principal investigator is publicly available through the University of North Carolina at Chapel Hill.

A.M. received personal fees and/or research grants from Novartis, Orion, Roche, personal fees from Servier, Adrenomed, from Abbott, Sanofi.

K.D. reports no conflicts.

N.J.S. reports no conflicts.

M.M. received consulting or speaker fees from Amgen, AstraZeneca, Bayer, Novartis, Relypsa, Servier, Stealth Therapeutics, Trevena, Abbott Vascular.

S.D.A. received grants from Vifor, Abbott Vascular, consultancy or speaking from Vifor, Bayer, Boehringer Ingelheim, Brahms, Janssen, Novartis, Servier, Stealth Peptides, and Astra.

J.G.C. reports no conflicts.

L.L.N. reports no conflicts.

F.Z. reports steering committee personal fees from Applied Therapeutics, Bayer, Boehringer,

Boston Scientific, Novartis, Janssen and CVRx, advisory board personal fees from, AstraZeneca,

Vifor Fresenius, Cardior, Cereno pharmaceutical and Merck, stock options at G3Pharmaceutical,

and being the founder of CardioRenal and CVCT.

G.S.F. has received committee fees and/or research grants from Novartis, Bayer, Vifor, Servier.

H.L.H. reports no conflicts.

P.P. received personal fees, research grants, and/or other from Novartis, Vifor, Amgen, Bayer,

BMS, Boehringer Ingelheim, Cardiorentis, Singulex, Fresenius, Cibiem.

D.J.v.V reports no conflicts.

C.C.L. received fees and/or research grants from Novartis, AstraZeneca, MSD.

P.v.d.M reports no conflicts.

J.N. received board speaker fees and/or other from Novartis, Roche Diagnostics, Abbott, Rovi, Vifor Pharma, Daiichi Sankyo, Boehringer Ingelheim, AstraZeneca.

A.B.-G. received board membership fees and/or other from Novartis, Roche Diagnostics, Vifor Pharma, Critical Diagnostics.

C.E. is an employee of Momentum Research who received research grants from Abbott Laboratories, Amgen, Celyad, Cirius Therapeutics, Roche Diagnostics, Sanofi, and Windtree Therapeutics Inc. A.A.V. received consultancy fees and/or research grants from Amgen, AstraZeneca, Bayer,
Boehringer Ingelheim, Cytokinetics, Merck, Novartis, Novonordisk, Roche diagnostics, Vifor .
G.C. is an employee of Momentum Research who received research grants from Abbott
Laboratories, Amgen, Celyad, Cirius Therapeutics, Roche Diagnostics, Sanofi, and Windtree
Therapeutics Inc.

#### **Author contributions**

The study was conceived by G.C. and designed by G.C., B.A.D., S.S., I.E.S., and A.A.V..

B.A.D., S.S., C.E., and I.E.S. analyzed the data.

G.C., B.A.D., and S.S. drafted the manuscript.

K.D., N.J.S., M.M., S.D.A., J.G.C., L.L.N., I.R.M., F.Z., G.S.F., H.L.H., P.P., D.J.V., C.C.L., P.M., J.N., A.B-G., and A.A.V. are members of the BIOSTAT-CHF consortium who acquired the data.

All authors critically reviewed the manuscript and contributed to the interpretation of results.

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#### **Figure Legends**

# Figure 1 (Central Illustration). Differential protein expression in inpatients relative to outpatients

Presented is a volcano plot of differential protein expression showing the fold change, i.e. the ratio of average expression in inpatients to average expression in outpatients, versus the corresponding t-test p-value per protein and on logarithmic scales. Higher values on the y-axis indicate stronger statistical significance, values >0 on the x-axis indicate upregulation in inpatients, and values <0 on the x-axis indicate downregulation in inpatients. Significantly differentially expressed proteins have been labeled.



index cohort Differential Protein Expression in in-hospital vs out-hospital patients

#### Table 1. Baseline characteristics by cohort and patient status

	Index	Cohort	Validatio	on Cohort
Parameter	Inpatient (N=1694)*	Outpatient (N=822)*	Inpatient (N=935)	Outpatient (N=803)
Male sex	1217 (71.8%)*	629 (76.5%)*	582 (62.2%)	563 (70.1%)
Age, years	70.4 (61.0, 78.3)	70.0 (61.5, 77.2)	76.1 (67.5, 82.7)	73.8 (66.3, 80.0)
Caucasian race	1671 (98.6%)	818 (99.5%)	930 (99.5%)	798 (99.4%)
Primary etiology: Ischemic heart disease	720 (43.2%)	406 (50.3%)	622 (66.5%)	500 (62.3%)
Primary etiology: Hypertension	166 (10.0%)	90 (11.2%)	116 (12.4%)	69 (8.6%)
Primary etiology: Cardiomyopathy	403 (24.2%)	228 (28.3%)	56 (6.0%)	67 (8.3%)
Primary etiology: Valvular disease	153 (9.2%)	37 (4.6%)	70 (7.5%)	68 (8.5%)
Reason for hospitalization				
New onset of Heart Failure	624 (36.8%)*	78 (9.5%)*		
HF hospitalization in last year	462 (27.3%)*	332 (40.4%)*	269 (29.3%)	191 (24.1%)
Atrial fibrillation	784 (46.3%)*	359 (43.7%)*	420 (45.3%)	340 (42.7%)
Myocardial infarction	605 (35.7%)	358 (43.6%)	469 (50.3%)	380 (47.4%)
PCI	341 (20.1%)	203 (24.7%)	179 (19.5%)	146 (18.3%)
CABG	289 (17.1%)	144 (17.5%)	154 (16.5%)	154 (19.2%)
Diabetes mellitus	583 (34.4%)*	236 (28.7%)*	326 (35.1%)	235 (29.4%)
COPD	304 (17.9%)*	132 (16.1%)*	206 (22.2%)	113 (14.2%)
Peripheral artery disease	204 (12.0%)	69 (8.4%)	179 (19.7%)	195 (24.8%)
Stroke	164 (9.7%)	69 (8.4%)	197 (21.2%)	118 (14.9%)
Hypertension	1064 (62.8%)	505 (61.4%)	572 (61.4%)	435 (54.4%)
Renal disease	520 (30.7%)	176 (21.4%)	453 (49.4%)	332 (41.8%)
Current malignancy	79 (4.7%)	18 (2.2%)	51 (5.5%)	28 (3.5%)
Pacemaker	127 (7.5%)	56 (6.8%)	57 (6.1%)	58 (7.2%)
ICD	135 (8.0%)	70 (8.5%)	24 (2.6%)	45 (5.6%)
Bi-ventricular pacer (CRT)	32 (1.9%)	17 (2.1%)	6 (0.6%)	21 (2.6%)
Bi-ventricular pacer (CRT) and ICD	102 (6.0%)	71 (8.6%)	16 (1.7%)	39 (4.9%)
Height, cm	170.0 (165.0, 177.0)	172.0 (166.0, 178.0)	168.0 (160.0, 176.0)	170.0 (161.0, 176.0)
Weight, kg	80.0 (69.0, 92.0)	81.0 (71.0, 91.0)	79.0 (67.0, 93.0)	82.0 (70.0, 95.0)
Body mass index, kg/m <sup>2</sup>	27.2 (24.0, 30.9)	27.1 (24.2, 30.4)	27.7 (24.1, 32.2)	28.6 (25.1, 32.9)
Heart rate, bpm	79.0 (70.0, 92.0)	72.0 (64.0, 82.0)	75.0 (64.0, 88.0)	69.0 (60.0, 79.0)

	Index	Cohort	Validation Cohort			
Parameter	Inpatient (N=1694)*	Outpatient (N=822)*	Inpatient (N=935)	Outpatient (N=803)		
Systolic blood pressure, mmHg	120.0 (110.0, 135.0)	126.0 (112.0, 140.0)	118.0 (106.0, 134.0)	129.0 (115.0, 145.0)		
Diastolic blood pressure	70.0 (64.0, 80.0)	80.0 (70.0, 85.0)	66.0 (58.0, 75.0)	71.0 (64.0, 79.0)		
LVEF, %	30.0 (25.0, 37.0)	30.0 (25.0, 35.0)	43.0 (35.0, 52.0)	39.0 (31.0, 48.0)		
Heart failure classification						
HFrEF (LVEF <40%)	1167 (79.0%)	652 (85.2%)	345 (41.4%)	381 (51.8%)		
HFmrEF (LVEF 40- <50%)	187 (12.7%)	95 (12.4%)	209 (25.1%)	192 (26.1%)		
HFpEF (LVEF ≥50%)	124 (8.4%)	18 (2.4%)	279 (33.5%)	162 (22.0%)		
NYHA class						
Class I	31 (1.9%)	25 (3.1%)	6 (0.6%)	11 (1.4%)		
Class II	456 (28.0%)	412 (50.6%)	235 (25.1%)	477 (59.5%)		
Class III	874 (53.6%)	354 (43.4%)	477 (51.0%)	295 (36.8%)		
Class IV	270 (16.6%)	24 (2.9%)	217 (23.2%)	19 (2.4%)		
Extent of peripheral edema						
Not Present	477 (33.2%)	366 (55.1%)	204 (24.2%)	385 (54.9%)		
Ankle	443 (30.9%)	189 (28.5%)	284 (33.7%)	199 (28.4%)		
Below Knee	375 (26.1%)	99 (14.9%)	278 (33.0%)	97 (13.8%)		
Above Knee	140 (9.8%)	10 (1.5%)	77 (9.1%)	20 (2.9%)		
Pulmonary congestion						
No	579 (34.8%)	575 (73.4%)	292 (32.3%)	627 (83.7%)		
Single base	219 (13.2%)	92 (11.7%)	59 (6.5%)	36 (4.8%)		
Bi-basilar	864 (52.0%)	116 (14.8%)	554 (61.2%)	86 (11.5%)		
Orthopnea	745 (44.1%)*	134 (16.3%)*				
Rales >1/3 up lung fields	229 (21.2%)	19 (9.1%)	43 (11.1%)	7 (3.3%)		
Elevated jugular venous pressure	435 (39.4%)*	119 (21.6%)*	281 (36.0%)	169 (24.5%)		
Hepatomegaly	270 (16.0%)	88 (10.8%)	34 (3.9%)	26 (3.8%)		
Aldosterone antagonists	913 (53.9%)	426 (51.8%)	274 (29.4%)	286 (35.8%)		
ACEi/ARB	1170 (69.1%)	650 (79.1%)	587 (63.0%)	632 (79.1%)		
Beta blocker	1356 (80.0%)	737 (89.7%)	644 (69.1%)	608 (76.1%)		
Loop diuretics	1684 (99.4%)	820 (99.8%)				
Diuretics	1692 (99.9%)	822 (100.0%)	920 (98.7%)	791 (99.0%)		
Digoxin	320 (18.9%)	171 (20.8%)	171 (18.3%)	138 (17.3%)		

	Index	Cohort	Validatio	n Cohort
Parameter	Inpatient (N=1694)*	Outpatient (N=822)*	Inpatient (N=935)	Outpatient (N=803)
Hemoglobin, g/dL	13.20 (11.76, 14.40)	13.60 (12.40, 14.70)	12.60 (11.20, 14.20)	13.50 (12.50, 14.70)
Hematocrit, %	40.00 (36.00, 43.20)	40.90 (37.60, 44.60)	39.70 (35.50, 43.70)	41.60 (38.50, 44.80)
Creatinine, µmol/L	104.31 (85.75, 132.60)	99.89 (81.00, 123.70)	99.00 (79.00, 128.00)	95.00 (81.00, 120.00)
eGFR, mL/min/1.73m <sup>2</sup>	58.60 (42.43, 75.79)	61.78 (46.95, 79.63)	57.91 (41.57, 77.35)	62.32 (45.82, 78.04)
Urea, mmol/L	12.10 (7.80, 19.50)	10.20 (7.40, 15.35)	9.00 (6.50, 12.90)	8.30 (6.50, 10.90)
Sodium, mmol/L	139.0 (137.0, 142.0)	140.0 (138.0, 142.0)	139.0 (136.0, 141.0)	140.0 (138.0, 142.0)
Potassium, mmol/L	4.20 (3.86, 4.56)	4.30 (4.02, 4.70)	4.20 (3.90, 4.50)	4.40 (4.10, 4.70)
Phosphate, mmol/L	0.87 (0.70, 1.05)	0.85 (0.69, 1.03)		
Albumin, g/L	32.0 (26.0, 36.0)	36.0 (30.0, 41.0)	35.0 (31.0, 39.0)	40.0 (37.0, 44.0)
Serum iron, µmol/L	7.0 (5.0, 11.0)	11.0 (7.0, 15.0)	9.0 (6.0, 13.0)	15.0 (11.0, 19.0)
Aldosterone, pg/mL	87.0 (38.0, 185.0)	109.0 (56.0, 215.0)		
White blood cell count, 10 <sup>9</sup> /L	7.90 (6.50, 9.80)	7.60 (6.30, 9.00)	7.70 (6.20, 9.60)	7.20 (6.00, 8.85)
Renin, μIU/mL	92.39 (28.40, 287.64)	78.66 (28.80, 215.46)		
Troponin T, pg/mL	36.40 (22.43, 62.34)	22.42 (14.53, 37.10)	43.81 (23.85, 99.00)	20.05 (13.46, 30.67)
Troponin I, pg/mL	15.59 (8.44, 36.61)	8.07 (4.84, 15.82)		
NT-proBNP, pg/mL	3291.5 (1422.5, 6880.5)	1921.5 (734.6, 3890.0)	2361.0 (952.0, 5851.0)	777.0 (309.0, 1768.0)
GDF-15, pg/mL	3130.0 (1951.0, 5299.0)	2115.5 (1335.0, 3337.0)	3481.0 (2185.5, 5818.5)	2272.0 (1570.0, 3549.0)
ET-1, pg/mL	5.68 (4.30, 7.53)	4.70 (3.62, 6.13)		
bio-ADM, pg/mL	36.89 (24.27, 59.50)	27.66 (19.95, 41.07)	32.1 (20.8, 52.3)	23.0 (16.6, 33.7)
IL-6, pg/mL	6.50 (3.65, 12.40)	3.20 (1.90, 5.70)		
CA125, U/mL	64.40 (20.60, 169.00)	19.80 (12.00, 43.80)	44.95 (20.30, 123.50)	17.50 (12.00, 29.00)

*Results are presented as frequency and percentage for discrete variables and as median, lower quartile (Q1), and upper quartile (Q3) for continuous variables.* \*Result presented in Ferreira JP, Metra M, Mordi I, Gregson J, Ter Maaten JM, Tromp J, Anker SD, Dickstein K, Hillege HL, Ng L, van Veldhuisen DJ, Lang CC, Voors AA, Zannad F. Heart failure in the outpatient versus inpatient setting: findings from the BIOSTAT-CHF study. Eur J Heart Fail. 2019;21(1):112-20.

Table 2. Ability of selected biomarkers to discriminate between patients with worsening heart failure enrolled in the inpatient versus the outpatient setting

Parameter*	Effect size for a change of	Odds ratio (95% CI)	P-value	Optimal cutpoint	AUC/ Observed C-index	Sensitivity	Specificit y	PPV	NPV
NT-proBNP, pg/mL, log2†	Doubling	1.32 (1.26, 1.39)	<.0001	9.2689 (616.90 pg/mL)	0.6395	0.9121	0.2182	0.7063	0.5466
NT-proBNP, log2 ≥9.2689‡	Yes vs. No	2.85 (2.24, 3.63)	<.0001	—	0.5640	0.9119	0.2162	0.7057	0.5435
NT-proBNP ≥400 pg/mL‡	Yes vs. No	2.64 (1.98, 3.51)	<.0001	_	0.5421	0.9400	0.1442	0.6936	0.5384
NT-proBNP (Olink)*	1	1.48 (1.37, 1.59)	<.0001	1.3156	0.6404	0.9175	0.2080	0.7049	0.5528
CA125, U/mL, log2†	Doubling	1.50 (1.42, 1.58)	<.0001	3.9078 (15.01 U/mL)	0.6983	0.8388	0.3680	0.7324	0.5256
CA125, log2 ≥3.9078‡	Yes vs. No	2.95 (2.40, 3.62)	<.0001	_	0.6017	0.8344	0.3690	0.7316	0.5195
CA125 ≥100 U/mL‡	Yes vs. No	4.25 (3.31, 5.46)	<.0001		0.6278	0.3832	0.8725	0.8610	0.4070
MUC-16 (Q8WXI7) †	1	1.58 (1.48, 1.69)	<.0001	4.8067	0.7118	0.8836	0.3028	0.7251	0.5666
ST2 (Q01638) †	1	1.84 (1.65, 2.04)	<.0001	2.4867	0.6629	0.9162	0.1861	0.700	0.5471
Troponin T, pg/mL, log2†	Doubling	1.80 (1.65, 1.96)	<.0001	3.8178 (14.10 pg/mL)	0.6837	0.9174	0.2436	0.7143	0.5889
Troponin T, log2 ≥3.8178‡	Yes vs. No	3.52 (2.76, 4.47)	<.0001	_	0.5794	0.9168	0.2420	0.7137	0.5852
Troponin I, pg/mL, log2†	Doubling	1.47 (1.38, 1.56)	<.0001	2.1205 (4.35 pg/mL)	0.6800	0.9198	0.2302	0.7120	0.6055
Troponin I, $\log 2 \ge 2.1205$ ;	Yes vs. No	2.91 (2.26, 3.76)	<.0001	_	0.5609	0.9204	0.2013	0.7037	0.5510
GDF-15, pg/mL, log2†	Doubling	1.72 (1.58, 1.88)	<.0001	10.0676 (1073.12 pg/mL)	0.6584	0.9443	0.1536	0.6969	0.5722
GDF-15, log2 ≥10.0676‡	Yes vs. No	2.83 (2.11, 3.80)	<.0001	_	0.5450	0.9423	0.1477	0.6950	0.5540
GDF-15 (Q99988) †	1	1.74 (1.58, 1.91)	<.0001	3.9365	0.6535	0.8966	0.2611	0.7147	0.5524
bio-ADM, pg/mL, log2†	Doubling	1.71 (1.55, 1.88)	<.0001	3.2218 (9.33 pg/mL)	0.6389	0.9893	0.0231	0.6762	n/a
bio-ADM, log2 ≥3.2218‡	Yes vs. No	1.43 (0.63, 3.22)	0.3914	—	0.5025	0.9885	0.0164	0.6744	0.4088

Parameter*	Effect size for a change of	Odds ratio (95% CI)	P-value	Optimal cutpoint	AUC/ Observed C-index	Sensitivity	Specificit y	PPV	NPV
ADM (P35318) †	1	1.48 (1.34, 1.63)	<.0001	1.3221	0.6377	0.9999	0.0004	0.6734	n/a
Results from logistic regression m	nodels, modeling	g the probability for in	ipatient.						
*Parameters followed by a UniPr	ot ID were meas	sured on Olink. The re	mainder were	e measured using Sir	ngulex assays.				
†Model parameters and optimal c	ut-point allowed	l to vary and averaged	l across impu	tation datasets.					
\$Same cut-point applied in each i	imputation datas	et and resulting param	neters average	ed across imputation	datasets.				

Table 3. Backward selection results Inpatient vs. Outpatient

		Multivariable Model 1 Local and central data		Multivariable Model 2 Local, central, and Singulex data		Multivariable Model 3 Local, central, Singulex, and Olink data	
Parameter	Effect size for a change of	Odds ratio (95% CI)	P-value	Odds ratio (95% CI)	P-value	Odds ratio (95% CI)	P-value
Age, years	5	0.93 (0.88, 0.97)	0.0017	0.94 (0.89, 0.98)	0.0100	0.89 (0.83, 0.95)	0.0008
LVEF, %	5	1.08 (1.03, 1.15)	0.0050	1.10 (1.04, 1.17)	0.0009	1.11 (1.04, 1.20)	0.0040
Diastolic blood pressure $\leq 70^{\dagger}$	5	0.70 (0.62, 0.79)	<.0001	0.69 (0.61, 0.78)	<.0001	0.74 (0.64, 0.85)	<.0001
Diastolic blood pressure $> 70^{\dagger}$	5	0.96 (0.90, 1.01)		0.97 (0.91, 1.03)		1.00 (0.93, 1.08)	
Heart rate, bpm	5	1.11 (1.08, 1.15)	<.0001	1.11 (1.07, 1.15)	<.0001	1.10 (1.05, 1.14)	<.0001
Albumin, g/L	1	0.88 (0.86, 0.91)	<.0001	0.92 (0.90, 0.94)	<.0001	0.93 (0.91, 0.96)	<.0001
Aldosterone, pg/mL, log2	Doubling	0.91 (0.86, 0.97)	0.0023				
ALT, U/L, log2	Doubling	1.25 (1.10, 1.42)	0.0007	1.28 (1.12, 1.46)	0.0004		
Hepcidin, nmol/L, log2	Doubling	1.17 (1.10, 1.24)	<.0001	1.09 (1.03, 1.16)	0.0041		
Serum iron, µmol/L, log2	Doubling	0.68 (0.58, 0.79)	<.0001	0.77 (0.66, 0.90)	0.0013		
Phosphate, mmol/L, log2	Doubling	3.42 (2.49, 4.69)	<.0001	3.30 (2.39, 4.56)	<.0001	2.28 (1.48, 3.49)	0.0002
sTfR, mg/L, log2	Doubling	1.38 (1.15, 1.65)	0.0005	1.48 (1.24, 1.77)	<.0001		
Total bilirubin, µmol/L, log2	Doubling	1.30 (1.10, 1.53)	0.0041				
Troponin T, pg/mL, log2	Doubling	1.44 (1.31, 1.60)	<.0001				
Potassium, mmol/L	1	0.68 (0.55, 0.83)	0.0002	0.67 (0.55, 0.82)	0.0002		
Transferrin, g/L	1	1.98 (1.45, 2.70)	<.0001				
Urea, mmol/L, log2	Doubling			1.31 (1.13, 1.53)	0.0007		
CA125, U/mL, log2	Doubling			1.25 (1.16, 1.35)	<.0001		
IL-6, pg/mL, log2	Doubling			1.26 (1.15, 1.38)	<.0001		
NT-proBNP, pg/mL, log2	Doubling			0.87 (0.81, 0.94)	0.0004	0.87 (0.78, 0.96)	0.0076
Troponin I, pg/mL, log2	Doubling			1.30 (1.21, 1.39)	<.0001	1.31 (1.20, 1.42)	<.0001
ADM (P35318)	1					0.70 (0.56, 0.87)	0.0014
AGRP (000253)	1					2.28 (1.65, 3.16)	<.0001
BOC (Q9BWV1)	1					0.45 (0.28, 0.73)	0.0012
FABP2 (P12104)	1					0.75 (0.64, 0.89)	0.0010
FGF-21 (Q9NSA1)	1					0.77 (0.70, 0.86)	<.0001

		X7 1/1 1 1 X7 1	14	N7 141 - 11 N7 1	10	Multivariable Mode	el 3	
		Multivariable Mode	el I ata	Local, central, and Singu	el 2 ilex data	Local, central, Singule Olink data	Local, central, Singulex, and Olink data	
Parameter	Effect size for a change of	Odds ratio (95% CI)	P-value	Odds ratio (95% CI)	P-value	Odds ratio (95% CI)	P-value	
GH (P01241)	1					0.75 (0.68, 0.83)	<.0001	
SERPINA12 (Q8IW75)	1					1.22 (1.06, 1.40)	0.0058	
FABP4 (P15090)	1					1.43 (1.22, 1.68)	<.0001	
FAS (P25445)	1					0.45 (0.32, 0.62)	<.0001	
GDF-15 (Q99988)	1					1.50 (1.15, 1.97)	0.0037	
MMP-3 (P08254)	1					0.57 (0.47, 0.69)	<.0001	
OPN (P10451)	1					1.55 (1.21, 1.98)	0.0006	
ST2 (Q01638)	1					1.66 (1.33, 2.09)	<.0001	
TR (P02786)	1					1.48 (1.18, 1.85)	0.0007	
GPNMB (Q14956)	1					0.25 (0.11, 0.56)	0.0015	
MUC-16 (Q8WXI7)	1					1.30 (1.17, 1.46)	<.0001	
RET (P07949)	1					0.53 (0.42, 0.66)	<.0001	
SYND1 (P18827)	1					2.13 (1.58, 2.86)	<.0001	
TFPI-2 (P48307)	1					1.90 (1.43, 2.53)	<.0001	
$uPA (P00749) \le 3.9^{+}$	1					0.31 (0.19, 0.48)	<.0001	
uPA (P00749) > 3.9 <sup>†</sup>	1					0.72 (0.46, 1.11)		
CYR61 (O00622) $\leq$ 7.2 <sup>†</sup>	1					1.64 (0.92, 2.92)	0.0006	
CYR61 (O00622) > 7.2 <sup>†</sup>	1					0.53 (0.38, 0.76)		
Hosmer-Lemeshow goodness-of-fit test			0.0741		0.0487		0.2342	
Pooled C-index in Index Cohort		0.8194		0.8324		0.9133		
	1 1		1	1	1		L	
C-index of final model in Index Cohort		0.8191		0.8322		0.9126		
Bias-corrected C-index (95% CI) §		0.8262 (0.8195, 0.8330)		0.8378 (0.8312, 0.8444)		0.9205 (0.9139, 0.9270)		
Bias-corrected difference in C-index (95% CI) compared to Model 1 <sup>§</sup>				0.0116 (0.0059, 0.0173)		0.0942 (0.0863, 0.1022)		

		Multivariable Model 1 Local and central data		Multivariable Model 2 Local, central, and Singulex data		Multivariable Model 3 Local, central, Singulex, and Olink data	
Parameter	Effect size for a change of	Odds ratio (95% CI)	P-value	Odds ratio (95% CI)	P-value	Odds ratio (95% CI)	P-value
Bias-corrected difference in C-index (95% CI) compared to Model 2 <sup>§</sup>						0.0827 (0.0744, 0.0909)	
C-index of final model in Validation Cohort		0.8731		0.8716		0.9011	
Results from logistic regression model,	, modeling the	e probability for inpatient.					
<sup>†</sup> Non-linear association modeled as lii <sup>§</sup> Bootstrap estimate with 100 resampli	near spline. ing steps.						

		Multivariable Model 1 Local and central data		Multivariable Model 2 Local, central, and Singulex data		Multivariable Model 3 Local, central, Singulex, and Olink data	
Parameter	Effect size for a change of	Hazard ratio (95% CI)	P-value	Hazard ratio (95% CI)	P-value	Hazard ratio (95% CI)	P-value
Age, years	5	1.15 (1.08, 1.24)	<.0001	1.11 (1.03, 1.19)	0.0042	1.13 (1.04, 1.22)	0.0031
Systolic blood pressure $\leq 125 \text{ mmHg}^{\dagger}$	10	0.83 (0.73, 0.95)	0.0075	0.85 (0.74, 0.97)	0.0301	0.83 (0.72, 0.95)	0.0017
Systolic blood pressure > 125 mmHg $^{\dagger}$	10	1.14 (1.03, 1.25)		1.11 (1.01, 1.23)		1.18 (1.07, 1.30)	
Albumin, g/L	1	0.97 (0.95, 0.99)	0.0007				
Hemoglobin, g/dL	1	0.92 (0.85, 1.00)	0.0389	0.91 (0.84, 0.99)	0.0245	0.86 (0.79, 0.94)	0.0009
FGF23, RU/mL, log2	Doubling	1.25 (1.15, 1.36)	<.0001				
NT-proBNP, pg/mL, log2	Doubling	1.31 (1.17, 1.46)	<.0001	1.20 (1.07, 1.35)	0.0019		
Renin, µIU/mL, log2	Doubling	1.17 (1.10, 1.24)	<.0001	1.16 (1.09, 1.24)	<.0001	1.14 (1.07, 1.23)	0.0002
Triglycerides, mmol/L, log2	Doubling	1.43 (1.02, 2.00)	0.0453				
Troponin T, pg/mL, log2	Doubling	1.18 (1.08, 1.30)	0.0004	1.17 (1.06, 1.29)	0.0015	1.14 (1.03, 1.26)	0.0088
Alkaline phosphatase, µg/L, log2	Doubling			1.30 (1.01, 1.67)	0.0452		
ET-1, pg/mL, log2	Doubling			1.49 (1.15, 1.93)	0.0024	1.45 (1.12, 1.87)	0.0044
Hepcidin, nmol/L, log2	Doubling			0.88 (0.82, 0.95)	0.0012	0.90 (0.84, 0.97)	0.0048
IL-6, pg/mL, log2	Doubling			1.11 (1.01, 1.23)	0.0323		
pro-ENK, pmol/L, log2	Doubling			1.45 (1.18, 1.77)	0.0003		
Transferrin, g/L	1			0.79 (0.64, 0.98)	0.0348		
Body mass index, kg/m <sup>2</sup>	1					0.96 (0.93, 0.99)	0.0072
HDL, mmol/L	1					0.47 (0.27, 0.81)	0.0084
LDL, mmol/L	1					1.27 (1.06, 1.53)	0.0158
BNP (P16860)	1					1.13 (1.02, 1.26)	0.0253
DCN (P07585)	1					2.10 (1.43, 3.08)	0.0002
TFF3 (Q07654)	1					1.21 (1.01, 1.45)	0.0422
IL6 (P05231) <sup>#</sup>	6.39 vs. 5.52					1.33 (1.15, 1.54)	<.0001
	5.52 vs. 4.83					1.10 (0.97, 1.24)	

#### Table 4. Backward selection results for Death through Day 180 - Inpatients

		Multivariable Model 1 Local and central data		Multivariable Model 2 Local, central, and Singulex data		Multivariable Model 3 Local, central, Singulex, and Olink data	
Parameter	Effect size for a change of	Hazard ratio (95% CI)	P-value	Hazard ratio (95% CI)	P-value	Hazard ratio (95% CI)	P-value
KRT19 (P08727) $\leq$ 3.2 <sup>†</sup>	1					0.67 (0.42, 1.08)	<.0001
KRT19 (P08727) > 3.2 <sup>†</sup>	1					1.67 (1.35, 2.06)	
Pooled C-index in Index Cohort		0.7973		0.7993		0.8281	
C-index of final model in Index Cohort		0.7969		0.7990		0.8280	
Bias-corrected C-index (95% CI) §		0.8090 (0.7942, 0.8238)		0.8114 (0.7965, 0.8263)		0.8485 (0.8336, 0.8634)	
Bias-corrected difference in C-index (95% CI) compared to Model 1 §				0.0024 (-0.0100, 0.0148)		0.0395 (0.0234, 0.0556)	
Bias-corrected difference in C-index (95% CI) compared to Model 2 <sup>§</sup>						0.0371 (0.0224, 0.0518)	
	•		•				-
Results from Cox proportional hazards <sup>†</sup> Non-linear association modeled as lin <sup>#</sup> Non-linear association modeled as cur <sup>§</sup> Bootstrap estimate with 100 resamplir	regression i ear spline. bic polynom ig steps.	model. iial. Effect sizes for 75th pe	rcentile v	s. median and for median vs	s. 25th per	rcentile are presented.	

Table 5. Backward selection	results for Death t	through Day 180 ·	- Outpatients
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		Multivariable Model 1 Local and central data		Multivariable Model 2 Local, central, and Singulex data		Multivariable Model 3 Local, central, Singulex, and Olink data	
Parameter	Effect size for a change of	Hazard ratio (95% CI)	P-value	Hazard ratio (95% CI)	P-value	Hazard ratio (95% CI)	P-value
Age, years	5	1.32 (1.09, 1.60)	0.0046	1.22 (1.02, 1.46)	0.0263	1.23 (1.02, 1.48)	0.0329
Hypertension	Yes vs. No	0.42 (0.20, 0.88)	0.0217				
Myocardial infarction	Yes vs. No	2.14 (1.08, 4.26)	0.0300				
Peripheral artery disease	Yes vs. No	3.73 (1.52, 9.15)	0.0044				
$LVEF \leq 35 \%^{\dagger}$	1	0.94 (0.89, 0.99)	0.0240				
LVEF > 35 % <sup>†</sup>	1	1.06 (1.00, 1.11)					
Hemoglobin, g/dL	1	0.75 (0.60, 0.94)	0.0130	0.72 (0.57, 0.90)	0.0039	0.72 (0.57, 0.91)	0.0056
Potassium, mmol/L	1	2.05 (1.04, 4.05)	0.0407				
Platelet count, 10 <sup>9</sup> /L, log2	Doubling	4.88 (2.25, 10.61)	0.0001	3.90 (1.54, 9.88)	0.0069	3.59 (1.31, 9.83)	0.0186
Total bilirubin, µmol/L, log2	Doubling	2.05 (1.23, 3.42)	0.0098				
Male sex	Yes vs. No			4.12 (1.51, 11.26)	0.0060	4.47 (1.57, 12.70)	0.0051
CA125, U/mL, log2	Doubling			1.47 (1.21, 1.78)	0.0001	1.45 (1.21, 1.74)	<.0001
pro-ENK, pmol/L, log2	Doubling			1.88 (1.12, 3.14)	0.0213		
Red blood cell count, 10 <sup>12</sup> /L	1			1.90 (1.10, 3.30)	0.0259	1.94 (1.11, 3.39)	0.0235
CNTN1 (Q12860)	1					0.53 (0.29, 0.96)	0.0409
SCGB3A2 (Q96PL1)	1					1.81 (1.31, 2.52)	0.0005
Pooled C-index in Index Cohort		0.8277		0.8156		0.8460	
C-index of final model in Index Cohort		0.8244		0.8127		0.8428	

Multivariable Model 1 Local and central data		el 1 ata	Multivariable Model 2 Local, central, and Singulex data		Multivariable Model 3 Local, central, Singulex, and Olink data		
Parameter	Effect size for a change of	Hazard ratio (95% CI)	P-value	Hazard ratio (95% CI)	P-value	Hazard ratio (95% CI)	P-value
Results from Cox proportional hazards <sup>†</sup> Non-linear association modeled as lir	regression i tear spline.	nodel.					

Table 6. Backward selection results for HF re-admission through Day 180 - Inpatients

	Multivariable Model 1 Local and central data		Multivariable Model 2 Local, central, and Singulex data		Multivariable Model 3 Local, central, Singulex, and Olink data		
Parameter	Effect size for a change of	Hazard ratio (95% CI)	P-value	Hazard ratio (95% CI)	P-value	Hazard ratio (95% CI)	P-value
Diabetes mellitus	Yes vs. No	1.33 (1.02, 1.74)	0.0373	1.33 (1.02, 1.74)	0.0373		
Myocardial infarction	Yes vs. No	1.75 (1.35, 2.27)	<.0001	1.75 (1.35, 2.27)	<.0001	1.87 (1.45, 2.41)	<.0001
Peripheral artery disease	Yes vs. No	0.66 (0.45, 0.98)	0.0386	0.66 (0.45, 0.98)	0.0386		
LVEF, %	1	1.02 (1.01, 1.03)	0.0043	1.02 (1.01, 1.03)	0.0043	1.02 (1.00, 1.03)	0.0118
Systolic blood pressure, mmHg	10	0.93 (0.87, 1.00)	0.0383	0.93 (0.87, 1.00)	0.0383	0.92 (0.87, 0.99)	0.0202
eGFR, mL/min/1.73m <sup>2</sup>	1	0.99 (0.99, 1.00)	0.0199	0.99 (0.99, 1.00)	0.0199		
Sodium, mmol/L	1	0.96 (0.93, 0.99)	0.0070	0.96 (0.93, 0.99)	0.0070	0.95 (0.92, 0.97)	0.0002
Alkaline phosphatase, µg/L, log2	Doubling	1.37 (1.09, 1.73)	0.0084	1.37 (1.09, 1.73)	0.0084	1.29 (1.02, 1.63)	0.0383
Ferritin, µg/L, log2	Doubling	0.89 (0.81, 0.99)	0.0241	0.89 (0.81, 0.99)	0.0241	0.89 (0.81, 0.98)	0.0136
FGF23, RU/mL, log2	Doubling	1.13 (1.02, 1.24)	0.0181	1.13 (1.02, 1.24)	0.0181		
GDF-15, pg/mL, log2	Doubling	0.80 (0.67, 0.95)	0.0135	0.80 (0.67, 0.95)	0.0135		
NT-proBNP, pg/mL, log2	Doubling	1.28 (1.16, 1.42)	<.0001	1.28 (1.16, 1.42)	<.0001		
Renin, µIU/mL, log2	Doubling	1.15 (1.08, 1.22)	<.0001	1.15 (1.08, 1.22)	<.0001		
HDL, mmol/L	1					0.62 (0.40, 0.97)	0.0401
RAGE (Q15109)	1					1.61 (1.20, 2.17)	0.0018
REN (P00797)	1					1.22 (1.05, 1.42)	0.0084
NT-proBNP (NA) $\leq$ 3.2 <sup>†</sup>	1					1.74 (1.30, 2.32)	0.0003
NT-proBNP (NA) $> 3.2^{\dagger}$	1					1.02 (0.88, 1.19)	
Pooled C-index in Index Cohort		0.7395		0.7395		0.7322	
C-index of final model in Index Cohort		0.7389		0.7389		0.7314	
Bias-corrected C-index (95% CI) <sup>§</sup>		0.7577 (0.7411, 0.7744)		0.7586 (0.7405, 0.7767)		0.7462 (0.7292, 0.7632)	

		Multivariable Model 1 Local and central data		Multivariable Model 2 Local, central, and Singulex data		Multivariable Model 3 Local, central, Singulex, and Olink data	
Parameter	Effect size for a change of	Hazard ratio (95% CI)	P-value	Hazard ratio (95% CI)	P-value	Hazard ratio (95% CI)	P-value
Bias-corrected difference in C-index (95% CI) compared to Model 1 <sup>§</sup>				0.0009 (-0.0072, 0.0089)		-0.0115 (-0.0308, 0.0078)	
Bias-corrected difference in C-index (95% CI) compared to Model 2 <sup>§</sup>						-0.0124 (-0.0310, 0.0062)	
Results from Cox proportional hazards <sup>†</sup> Non-linear association modeled as lin <sup>§</sup> Bootstrap estimate with 100 resampli	regression 1 pear spline. ng steps.	model.					

Table 7.	Backward	selection	results for	HF re-	-admission	through	Dav	180 - (	Outpati	ients
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		Multivariable Model 1 Local and central data		Multivariable Model 2 Local, central, and Singulex data		Multivariable Model 3 Local, central, Singulex, and Olink data		
Parameter	Effect size for a change of	Hazard ratio (95% CI)	P-value	Hazard ratio (95% CI)	P-value	Hazard ratio (95% CI)	P-value	
HF hospitalization in last year	Yes vs. No	1.69 (1.05, 2.72)	0.0318	1.85 (1.15, 2.99)	0.0119	1.77 (1.09, 2.87)	0.0202	
Body mass index $\leq 25 \text{ kg/m}^{2 \dagger}$	1	0.84 (0.74, 0.96)	0.0334	0.81 (0.71, 0.92)	0.0050	0.82 (0.72, 0.94)	0.0212	
Body mass index > 25 kg/m <sup>2</sup> <sup>†</sup>	1	1.03 (0.97, 1.09)		1.02 (0.96, 1.09)		1.04 (0.98, 1.11)		
LVEF, %	1	1.03 (1.01, 1.05)	0.0066	1.03 (1.01, 1.06)	0.0016	1.05 (1.03, 1.08)	<.0001	
Diastolic blood pressure	5	0.84 (0.76, 0.94)	0.0015	0.84 (0.75, 0.94)	0.0019			
Alkaline phosphatase, µg/L, log2	Doubling	1.55 (1.01, 2.37)	0.0481	1.57 (1.02, 2.41)	0.0428			
FGF23, RU/mL, log2	Doubling	1.32 (1.14, 1.53)	0.0003	1.29 (1.11, 1.49)	0.0008			
NT-proBNP, pg/mL, log2	Doubling	1.25 (1.06, 1.48)	0.0084					
ET-1, pg/mL, log2	Doubling			1.83 (1.17, 2.85)	0.0081			
Troponin I, pg/mL, log2	Doubling			1.16 (1.02, 1.33)	0.0285			
Systolic blood pressure, mmHg	10					0.82 (0.71, 0.94)	0.0040	
Serum iron, µmol/L, log2	Doubling					0.73 (0.58, 0.94)	0.0138	
ADAM-TS13 (Q76LX8)	1					0.33 (0.11, 0.95)	0.0411	
VEGFD (O43915)	1					3.04 (1.62, 5.72)	0.0006	
NT-proBNP (NA)	1					1.55 (1.27, 1.89)	<.0001	
GIF (P27352) $\leq$ 9.4 <sup>†</sup>	1					0.81 (0.65, 1.02)	0.0060	
GIF (P27352) > 9.4 <sup>†</sup>	1					2.33 (1.37, 3.97)		
Pooled C-index in Index Cohort		0.7966		0.7984		0.8234		
C-index of final model in Index Cohort		0.7960		0.7981		0.8231		
Results from Cox proportional hazards to <sup>†</sup> Non-linear association modeled as line	regression 1 ear spline.	nodel.						